

Problem Set 3 - Solutions

(In the below, I just show the expected math. In your work you should give justification, often from the notes, for each step.)

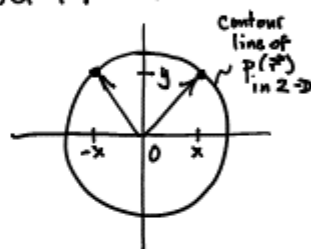
1.

a) Write \vec{r} as (x, y, z) , & $\nabla P_t(\vec{R})$ as (a, b, c) (as \vec{r} -indep. vector).

$$\begin{aligned} & \int \rho(\vec{r}) (\vec{r} \cdot \nabla) P_t(\vec{R}) d^3 \vec{r} \\ &= \int \rho(\vec{r}) [(x, y, z) \cdot (a, b, c)] d^3 \vec{r} \\ &= a \int x \rho(\vec{r}) d^3 \vec{r} + b \int y \rho(\vec{r}) d^3 \vec{r} + c \int z \rho(\vec{r}) d^3 \vec{r}. \quad (1) \end{aligned}$$

Observe $\rho((-x, y, z)) = \rho((x, y, z))$. So for y & z fixed x is odd & $\rho((x, y, z))$ is even, so the integral over x of $x \rho((x, y, z))$ is zero. Hence

$$\int x \rho(\vec{r}) d^3 \vec{r} = 0$$



By analogous reasoning, all 3 integrals in (1) are zero, so (1) is zero.

b)
$$\begin{aligned} & \int \rho(\vec{r}) (\vec{r} \cdot \nabla)^2 P_t(\vec{R}) d^3 \vec{r} \\ &= \int \rho(\vec{r}) [(x, y, z) \cdot (\frac{\partial}{\partial x}, \frac{\partial}{\partial y}, \frac{\partial}{\partial z})]^2 P_t(\vec{R}) d^3 \vec{r} \\ &= \int \rho(\vec{r}) (x \frac{\partial}{\partial x} + y \frac{\partial}{\partial y} + z \frac{\partial}{\partial z})^2 P_t(\vec{R}) d^3 \vec{r} \quad \sim (2) \\ &= \int \rho(\vec{r}) (x^2 \frac{\partial^2}{\partial x^2} + y^2 \frac{\partial^2}{\partial y^2} + z^2 \frac{\partial^2}{\partial z^2} + 2xy \frac{\partial^2}{\partial x \partial y} + 2xz \frac{\partial^2}{\partial x \partial z} + 2yz \frac{\partial^2}{\partial y \partial z}) P_t(\vec{R}) d^3 \vec{r} \end{aligned}$$

By reasoning analogous to that above, all the cross terms such as $\int xy \rho(\vec{r}) d^3 \vec{r}$ are zero. So (2) reduces to

$$\begin{aligned} & \int x^2 \rho(\vec{r}) d^3 \vec{r} \frac{\partial^2}{\partial x^2} P_t(\vec{R}) + \int y^2 \rho(\vec{r}) d^3 \vec{r} \frac{\partial^2}{\partial y^2} P_t(\vec{R}) + \int z^2 \rho(\vec{r}) d^3 \vec{r} \frac{\partial^2}{\partial z^2} P_t(\vec{R}) \\ &= \langle x^2 \rangle \nabla^2 P_t(\vec{R}) \quad (\text{by the isotropy of } \rho(\vec{r}), \langle x^2 \rangle = \langle y^2 \rangle = \langle z^2 \rangle) \\ &= \frac{3 \langle x^2 \rangle}{3} \nabla^2 P_t(\vec{R}) \\ &= \frac{\langle \vec{r}^2 \rangle}{3} \nabla^2 P_t(\vec{R}) \end{aligned}$$

2.

a)
$$E[e^{itX}] = \frac{1}{2} e^{it(1)} + \frac{1}{2} e^{it(-1)} = \cos t$$

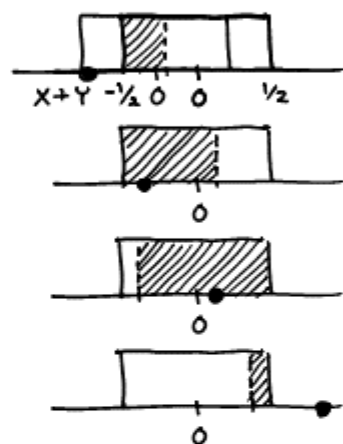
$$\begin{aligned}
 \text{b) } E[e^{itX}] &= \int_{-c}^c \frac{1}{2c} e^{itx} dx = \int_{-c}^c \frac{1}{2c} \cos(tx) dx + i \int_{-c}^c \frac{1}{2c} \sin(tx) dx \\
 &= \frac{1}{2c} \frac{\sin(tx)}{t} \Big|_{-c}^c = \frac{\sin(ct) - \sin(-ct)}{2ct} = \frac{\sin(ct)}{ct}
 \end{aligned}$$

0, by odd symmetry

c) Observe that the sum of 2 i.i.d. r.v.s, $X+Y$, each w/ uniform dist. on $(-1/2, 1/2)$, is a r.v. w/ a triangular distribution $1-|x|$, where $x \in (-1, 1)$. We see this as follows (see below): We move the final value of $X+Y$ from left to right & integrate over the shaded region, which corresponds to the interval of values of X that can yield the specified value of $X+Y$. Computed pointwise, the value of this integral is clearly $1-|x|$, $x \in (-1, 1)$.

Hence, the ch.f. of the triangular dist. is:

$$\begin{aligned}
 E[e^{it(X+Y)}] &= E[e^{itX}] E[e^{itY}] \\
 &= \left(\frac{\sin(t/2)}{t/2} \right)^2 \\
 &= \frac{2(1-\cos t)}{t^2} \quad (\text{since } \cos^2 \theta = 1 - 2\sin^2 \theta)
 \end{aligned}$$



3.

a) The Cauchy distribution is a stable law w/ $\alpha=1$ & $\kappa=0$.

So the ch.f. of X_k is $E[e^{itX_k}] = \exp(itc - b|t|)$ where c, b are constants. Therefore, the ch.f. of $\frac{1}{n}(X_1 + X_2 + \dots + X_n)$ is

$$\begin{aligned}
 &E[e^{i(t/n)(X_1 + X_2 + \dots + X_n)}] \\
 &= \prod_{k=1}^n E[e^{i(t/n)X_k}] = \\
 &= \prod_{k=1}^n \exp(i(t/n)c - b|t/n|) \\
 &= \exp\left(i \sum_{k=1}^n (t/n)c - b \sum_{k=1}^n |t/n|\right) \\
 &= \exp(itc - b|t|) = E[e^{itX_1}].
 \end{aligned}$$

Since $\frac{1}{n}(X_1 + X_2 + \dots + X_n)$ & X_1 have the same ch.f. they have the same distribution.

b) Take x to be positive.

$$\Pr[1/X_k > x] = \Pr[X_k < x^{-1}]$$

As we take $x \rightarrow \infty$, the rectangular area at right becomes an increasingly better approximation of $\Pr[X_k < x^{-1}]$, i.e. $\Pr[1/X_k > x] \sim f(0)/x$ as $x \rightarrow \infty$.

By the same argument reflected across 0,

$$\Pr[1/X_k < -x] \sim f(0)/x \text{ as } x \rightarrow \infty.$$

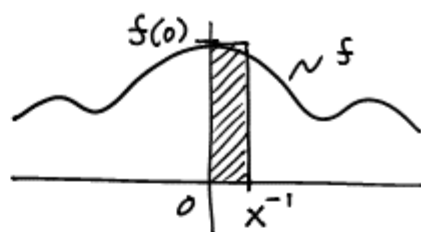
Hence,

$$\begin{aligned} \text{(i.) } \lim_{x \rightarrow \infty} \frac{\Pr[1/X_k > x]}{\Pr[|1/X_k| > x]} &= \lim_{x \rightarrow \infty} \frac{\Pr[1/X_k > x]}{\Pr[1/X_k > x] + \Pr[1/X_k < -x]} \\ &= \lim_{x \rightarrow \infty} \frac{f(0)/x}{f(0)/x + f(0)/x} = \frac{1}{2} \in [0, 1] \end{aligned}$$

$$\begin{aligned} \text{(ii.) } \Pr[|1/X_k| > x] &= \Pr[1/X_k > x] + \Pr[1/X_k < -x] \sim x^{-1}f(0) + x^{-1}f(0) \\ &= x^{-1}2f(0) \end{aligned}$$

$2f(0)$ is trivially slowly varying ($\lim_{x \rightarrow \infty} \frac{2f(0)}{2f(0)} = 1$).

The above show that the mean $\frac{1}{n}(\frac{1}{x_1} + \frac{1}{x_2} + \dots + \frac{1}{x_n})$ satisfies the condition of the thm on slides 14-15 of Lecture 6, w/ $\theta = \frac{1}{2}$ & $\alpha = 1$. So $k = 2\theta - 1 = 0$, & by case (b) on slide 15, the mean is a Cauchy r.v.



4. (***) is $U_k(x) = p x U_{k+1}(x) + q x U_{k-1}(x) \quad 1 < k < N$

For simplicity, consider just one x .

① First, by substitution of the form,

$$U_k(x) = A(x) \lambda_+^k(x) + B(x) \lambda_-^k(x) \quad (\#)$$

into (**), & the def. of $\lambda_{\pm}^k(x)$, this form does satisfy (**), so the $U_k(x)$ do provide a soln to (**).

② Second, for arbitrary initial values c_0 & c_1 , we can set $A(x)$ & $B(x)$ in (#) such that $U_0(x) = c_0$ & $U_1(x) = c_1$, since this is just a 2-D linear problem.

③ Third, let $V_k(x)$ be an arbitrary soln satisfying (**), & choose c_0 & c_1 to be $V_0(x)$ & $V_1(x)$, respectively. Then by (**), $V_2(x)$ must be just the same as the $U_2(x)$ implied by (†) w/ $A(x)$ & $B(x)$ as found in part ②. Likewise, $V_3(x)$ must be the same as $U_3(x)$, & so on. So $V_k(x) = U_k(x)$ for all k , i.e. all $V_k(x)$ must have the form (†).

5. First, we evaluate $\lambda_{\pm}(1)$ 2 different ways:

$$\begin{aligned}\lambda_{\pm}(1) &= \frac{1 \pm \sqrt{1-4pq}}{2p} = \frac{1 \pm \sqrt{1-4p(1-p)}}{2p} = \frac{1 \pm \sqrt{(1-2p)^2}}{2p} = \frac{1 \pm |1-2p|}{2p} \\ &= \frac{1 \pm \sqrt{1-4(1-q)q}}{2p} = \frac{1 \pm \sqrt{(1-2q)^2}}{2p} = \frac{1 \pm |1-2q|}{2p}\end{aligned}$$

Suppose $p = \frac{1}{4}$. Then the above indicates $\lambda_+(1) = 2(1 + \frac{1}{2}) = 3$
& $\lambda_-(1) = 1$

By the def. of U_k , $U_k(x) \leq U_k(1) \leq 1$, where the 2nd inequality holds since $U_k(1)$ is the prob. of going bankrupt starting from k . For $p = \frac{1}{4}$,
 $U_k(x) = A(x)3^k + B(x)$.

So for any $A(x) \neq 0$, there is some k large enough that
 $U_k(x) \notin [0, 1]$,

which is not allowed for a probability. So, $A(x) = 0$ for all x . Hence,
 $U_k(x) = B(x)\lambda_{-}^k(x)$.

But the initial condition $U_0(x) = 1$ still holds, so $B(x) = 1$ &
 $U_k(x) = \lambda_{-}^k(x)$.

The above evaluation of $\lambda_{\pm}(1)$ gives that if $p \leq q$, then

$$\lambda_{-}^k(1) = \frac{1 - (1-2p)}{2p} = 1,$$

and if $p \geq q$, then

$$\lambda_{-}^k(1) = \frac{1 - (1-2q)}{2p} = \frac{q}{p}.$$

So, $U_k(1) = \begin{cases} 1 & p \geq q \\ (q/p)^k & p < q \end{cases}$.